Statistical Techniques II

EXST7015

Analysis of Variance



15a_ANOVA_Introduction 1

Design

- The simplest model for Analysis of Variance (ANOVA) is the CRD, the Completely Randomized Design
- This model is also known a "One-way" Analysis of Variance.
- Unlike regression, which fits slopes for regression lines and calculates a measure of random variation about those lines, ANOVA fits means and variation about those means.

Design (continued)

The hypotheses tested are hypotheses about the equality of means

•
$$H_0: \mu_1 = \mu_2 = \mu_3 = \mu_4 = \dots = \mu_t$$

► Where

- the μ_i represent means of the levels of some categorical variable
- "t" is the number of levels in the categorical variable.
- H₁: some µ_i is different

Design (continued)

- We will generically refer to the categorical variable as the "treatment" even though it may not actually be an experimenter manipulated effect.
 - The number of treatments will be designated "t".
 - The number of observations within treatments will be designated n for a balanced design (the same number of observations in each treatment), or n_i for an unbalanced design (for i = 1 to t).

Design (continued)

- The assumptions for basic ANOVA are very similar to those of regression.
 - The residuals, or deviations of observations within groups should be normally distributed.
 - The treatments are independently sampled.
 - The variance of each treatment is the same (homogeneous variance).

ANOVA review

- I am borrowing some material from my EXST7005 notes on t-test and ANOVA.
 See those notes for a more complete review of the introduction to Analysis of Variance (ANOVA).
- Start with the logic behind ANOVA.

- Prior to R. A. Fisher's development of ANOVA, investigators were likely to have used a series of t tests to test among t treatment levels.
- What is wrong with that? Recall the Bonferroni adjustment. Each time we do a test we increase the chance of error. To test among 3 treatments we need to do 3 tests, among 4 treatments, 6 tests, 5 treatments are 10 tests, etc.

- What is needed is ONE test for a difference among all tests with one overall value of α specified by the investigator (usually 0.05).
- Fisher's solution was simple, but elegant.

Suppose we have a treatment with 5 categories or levels. We can calculate a mean and variance for each treatment level. In order to get one really good estimate of variance we can pool the individual variances of the 5 categories (assuming homogeneity of variance). This pooled variance can be calculated as a weighted mean of the variance (weighted by the degrees of freedom).



And since (n₁-1)S₁² = SS₁, the weighted mean is simply the sum of the SS divided by the sum of the d.f.

$$S_{p}^{2} = \frac{(n_{1} - 1)S_{1}^{2} + (n_{2} - 1)S_{2}^{2} + (n_{3} - 1)S_{3}^{2} + (n_{4} - 1)S_{4}^{2} + (n_{5} - 1)S_{5}^{2}}{(n_{1} - 1) + (n_{2} - 1) + (n_{3} - 1) + (n_{4} - 1) + (n_{5} - 1)}$$

$$S_p^2 = \frac{SS_1 + SS_2 + SS_3 + SS_4 + SS_5}{(n_1 - 1) + (n_2 - 1) + (n_3 - 1) + (n_4 - 1) + (n_5 - 1)}$$

R. A. Fisher 1890 - 1962



1913



1952

15a_ANOVA_Introduction 12

- So we have one very good estimate of the random variation, or sampling error, S².
- Then what?

R. A. Fisher 1929



- Now consider the treatments. Why don't they all fall on the overall mean?
- Actually, under the null hypothesis, they should, except for some random variation.
- So if we estimate that random variation, it should be equal to the same error we already estimated within groups?

If we estimate a variance with means, we are estimating the variance of means, which is S²/n. If we multiply this by "n" it should actually be equal to S², which we estimated with S²_p, the pooled variance estimate.

So if the null hypothesis is true, the mean square of the deviations within groups should be equal to the mean square of the deviations of the means multiplied by "n"!!!!



15a_ANOVA_Introduction 17

ANOVA review (continued) Now, if the null hypothesis is not true, and some μ_i is different, then what?



- Then, when we calculate a mean square of deviations of the means from the overall mean, it should be larger than the previously estimated S²_p.
- So we have two estimates of variance, S²_p and the variance from the treatment means. If the null hypothesis is true, they should not be significantly different.

If the null hypothesis is FALSE, the treatment mean square should be larger. It will therefore be a ONE TAILED TEST!

R. A. Fisher 1936



We usually present this in an "Analysis of Variance" table.

| Source | d.f. | Sum of Squares | Mean Square |
|-----------|--------|-------------------|-------------|
| Treatment | t-1 | SSTreatment | MSTreatment |
| Error | t(n-1) | SSError | MSError |
| Total | tn-1 | SSTotal | |

Degrees of freedom

- There are tn observations total (Σn_i if unbalanced).
- After the correction factor, there are tn-1 for the corrected total.
- There are t-1 degrees of freedom for the t treatment levels.
- Each group contributes n-1 d.f. to the pooled error term. There are t groups, so the pooled error (MSE) has t(n-1) d.f.

- The SSTreatments is the SS deviations of the treatment means from the overall mean.
- Each deviation is denoted \u03c6_i, and is called an treatment "effect".

SSTreatments =
$$\sum_{i=1}^{t} (\overline{Y}_i - \overline{\overline{Y}})^2 = \sum_{i=1}^{t} \mathcal{T}_i^2$$

The model for regression is

$$\mathbf{P} \mathbf{Y}_{i} = \beta_{0} + \beta_{1} \mathbf{X}_{i} + \varepsilon_{i}$$

The effects model for a CRD is

$$\blacktriangleright \mathbf{Y}_{ij} = \boldsymbol{\mu} + \boldsymbol{\tau}_i + \boldsymbol{\varepsilon}_{ij}$$

- -where the treatments are i=1, 2, ... t
- and the observations are j=1, 2, ... n, or n_i
 for unbalanced data
- The means model is $Y_{ij} = \mu_i + \varepsilon_{ij}$

- The calculations.
- The SSTotal is exactly the same as regression, the sum of all ΣY²_{ij} observations (squared first).
- The correction factor is exactly the same too, all observations are summed, the sum is squared and divided by the number of observations, (ΣY_{ij})²/tn.

n

1=1

1 = 1

UncorrectedSSTreatments = $\sum (\sum Y_{ij})$

| Obs | Group | Group | Group | Group |
|------|-----------------------------|-----------------------------|-----------------------------|-----------------------------|
| | 1 | 2 | 3 | 4 |
| 1 | Y ₁₁ | Y ₂₁ | Y ₃₁ | Y ₄₁ |
| 2 | Y ₁₂ | Y ₂₂ | Y ₃₂ | Y ₄₂ |
| | | | | |
| n | Y _{1n} | Y _{2n} | Y _{3n} | Y _{4n} |
| sum | $\Sigma \mathbf{Y}_{1}$ | $\Sigma \mathbf{Y}_2$ | $\Sigma \mathbf{Y}_{3}$ | $\Sigma \mathbf{Y}_{4}$ |
| mean | $\overline{\mathbf{Y}}_{1}$ | $\overline{\mathbf{Y}}_{2}$ | $\overline{\mathbf{Y}}_{3}$ | $\overline{\mathbf{Y}}_{4}$ |
| | | | | |

- Calculations are the same as regression for the corrected sum of squares total.
- The corrected SS treatments is the uncorrected treatments calculated from the marginals less the same correction factor used for the total.
- Error is usually calculated as the SSTotal minus the SSTreatments.

- We use an F test to test the equality of two means. An Analysis of Variance usually proceeds with an F test of the MSTreatment using the MSError. The test has t-1 and t(n-1) degrees of freedom.
- This F test will be ONE TAILED since we expect the treatment variance to be too large if the null hypothesis is not true.

- The MSError estimate a variance we designate σ² or σ²_ε.
- If the null hypothesis is true, the MSTreatments estimates the SAME VARIANCE, σ².
- However, if the null hypothesis is false the MSTreatment variance is the same σ^2 plus some amount due to the differences between treatments. This is designated $\sigma^2 + n\sigma^2_{\tau}$.

- Since the treatment variance can be designated $\sigma^2 + n\sigma^2_{\tau}$, we can see that the null hypothesis can be stated as either the usual H₀: $\mu_1=\mu_2=...=\mu_t$, or as H₀: $\sigma^2_{\tau}=0$.
- Which is best depends on the nature of the treatment. If the treatment levels are randomly chosen from a large number of treatment levels, then they estimate the variance of that treatment population and would be random. This would be σ²_τ.

However, if the treatments are not chosen from a large number of treatments; if they are either all of the levels of interest or all of the levels that exist, then they are said to be FIXED. Fixed treatment levels represent a group of means that are of interest to the investigator, so $H_0:\mu_1=\mu_2=...=\mu_t$ is a better representation of the null hypothesis than $H_0:\sigma^2_{\tau}=0$.

- For fixed treatments we still calculate a sum of squared treatment effects and divide by d.f. This is designated
 - $n\Sigma \tau_{i}^{2}$,/(n-1) and the F test is the same. These simply do not represent a variance.
- The two values estimated by the MSTreatment (σ² + nσ², or σ² + nΣτ²,/(n-1)) and MSError (σ²) are called <u>expected</u> <u>mean squares</u>.

- One final note on the F test.
- Given that MSTreatments and MSError estimate these EMS (expected mean squares) we can rewrite the F test as F=(σ² + nσ²_τ)/(σ²).
- From this value we can see that it must be a one tailed test because nσ², cannot be negative, so the ratio is always >1. We can also see that increasing n increases power.

Example

See SAS handout.

R. A. Fisher 1946



15a_ANOVA_Introduction 34



15a_ANOVA_Introduction 35